family size) and the oldest women (who had limited access to contraception in their early childbearing years) were more likely than those aged 25–39 to have never practiced contraception (unadjusted odds ratios of 2.1 and 3.3, respectively). The trend was also U-shaped rather than linear for the association between the woman’s number of live births and her likelihood of nonuse. Women who had had at least six children were almost 3.7 times as likely to have never used contraception, while those who had had only 0–1 live births were 2.4 times as likely to have never practiced contraception.

The bivariate analysis uncovered no significant association between the likelihood of nonuse and the distance from the woman’s residence to the nearest medical clinic, however, or between nonuse and age at first union.

Multivariate Analyses
Since the bivariate analysis suggested that many other background factors besides education might explain the likelihood of nonuse, we controlled for all significant variables in a multivariate analysis (Table 3). Seven of the 12 variables that were significant in the bivariate analysis retained their significance in the multivariate analysis. All seven explanatory variables independently raised the probability of nonuse by factors ranging from 1.6 to 2.8. These were, in the order of the magnitude of the odds ratios, the number of children who died (in the comparison between more than two deaths and no deaths only), the age at first pregnancy, family structure, the place of the most recent delivery, paid employment, type of flooring and education—but only in the comparison between illiterate women and those who had completed a secondary education. (As in the bivariate analysis, the multivariate analysis revealed no significant difference in the likelihood of nonuse when women who had some primary school or had finished primary school were compared with those who had had at least a secondary education.)

As the multivariate model suggested that schooling affected nonuse indirectly, we calculated Spearman rank correlation coefficients to identify those explanatory variables through which education would act. Schooling was significantly correlated with all of the other explanatory variables for nonuse, especially the place of last delivery (rho=0.2618, p<.0001), the number of children who died (rho=-0.2505, p<.0001) and age at first pregnancy (rho=0.2014, p<.0001). The association between education and family structure was slightly more strong (rho=0.1434, p<.0001), while that between education and current employment was relatively weak (rho=-0.0826, p=.0154).

Although we considered age in our initial multivariate model, the fact that women have different contraceptive needs and desires over the course of the reproductive life cycle led us to conduct additional multivariate analyses among women of three distinct age-groups—15–24-year-olds, 25–39-year-olds and 40–49-year-olds. Further, because such a large proportion of the entire sample was younger than age 40 (81%), these initial results were clearly biased toward explanations that relate primarily to women younger than 40 years of age.

In the separate multivariate analyses conducted with 40–49-year-olds (Table 4, page 136), no schooling at all (p=.0036) emerged as a strong predictor of nonuse: Among these oldest women, those who never attended school were more than three times as likely as those who attended secondary school to have never practiced contraception (adjusted odds ratio of 3.1).* The other variables that independently affected the likelihood of never-use among the oldest women were where they last delivered (with those who delivered

*The analysis also showed no significant association when older women with incomplete primary school were compared with those who had completed secondary school. However, the association was significant, but unexpected and in the opposite direction, when the likelihood of nonuse among older women who had finished primary school was compared with that among older women who had finished secondary school (adjusted odds ratio of 0.10). We believe that this inconsistency can be explained by the interaction between the schooling, family structure and nonuse variables, which we evaluated using a hierarchical log linear model (with three degrees of freedom, χ²= 8.749, p=.0238).